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**Modelling the age-dependent personal  
income distribution in the USA**

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## **Modelling the age-dependent personal income distribution in the USA**

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### **Abstract**

Numerical modelling of the age-dependent personal income distribution (PID) in the USA is fulfilled based on a micro- and macroeconomic model and results of the overall PID modelling. As expected from the age-dependent accuracy of the Current Population Surveys, the model has demonstrated an excellent prediction power in almost all income bins except the lowermost ones.

Here we address the problem of the fine age structure of the PIDs. The age-dependent PIDs are modelled by using the same defining parameters as the overall PIDs. The predicted PIDs accurately describe the observed ones reproducing such complex features as the exponential PID decay in the youngest and oldest age groups. The evolution of the age-dependent PIDs in time is also accurately predicted. The difference in the PID levels in the youngest age group is explained by some shortcomings in the design of the enumeration procedure. Corresponding recommendations are given in order to improve the PID estimates.

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## **Introduction**

Numerical modelling of the age-dependent personal income distribution (PID) in the USA is fulfilled based on a micro- and macroeconomic model and results of the overall PID modelling. Kitov (2005a) formally introduced a microeconomic model for a personal income evolution as a function of individual capacity to earn money and economic growth. The sum of all the personal incomes predicted by the microeconomic model (with a simple assumption about the distribution of the capacity to earn money) builds a macroeconomic model. Defining parameters of the macroeconomic model providing the best fit between the observed and predicted PIDs are presented and briefly discussed in the paper.

Kitov (2005b) presented results of the overall PID modelling. Principal features of the observed PIDs' are revealed and discussed. The uncovered characteristics provide a deterministic description of the PID shape and its evolution in time. As a result, the macroeconomic model accurately describes the observed evolution of the overall PID between 1994 and 2002. The two papers present a concept of two branches of the PIDs: a lower income branch and a high income branch. The former is accurately described by the model and the latter is a standard Pareto distribution. The Pareto distribution (or power law distribution) results from a number of processes called as a whole "self-organized criticality". There is no economic model available to formally express the processes leading to the observed Pareto distribution and only the number of people governed by this distribution is predicted by the developed macroeconomic model. The number of people is the only parameter needed to model the observed Pareto distribution because other parameters of the distribution are fixed for a given society.

The overall PIDs presented in (Kitov 2005b) include the personal incomes of all the Americans above 15 years of age as published by the U.S. Census Bureau (2004a). By design, aggregation in narrow income bins of \$2500 does not depend on age. It is obvious, however, that each and every personal income undergoes important changes with time. The starting point for any personal income is apparently a zero income. Then the income grows with decreasing rate to some maximum value. Between 1994 and 2002, the observed personal incomes in the USA usually reached their relative maximum value at some age between 45 and 55 years and then dropped. The overall PIDs are not able to describe all these complex features and processes due to lack of age resolution. The processes and features, however, are extremely important from a personal point of view as a prospective of income trajectory. They are also a big challenge to any model for the personal income distribution and its evolution. In Section 1 of this paper, we briefly discuss some problems related to the PID and its evolution in various age groups. The high resolution data set under analysis is obtained from the same source as in the previous papers – the U.S. Census Bureau (2004). In Section 2, we consistently apply the microeconomic and macroeconomic model to accurately predict the observed behaviour of the age-dependent PIDs. Some principal results of the modelling are presented in graphic form. Conclusions made in Section 3 are discussed in a broader economic context.

### **1. Evolution of the personal income distribution in various age groups: Observations**

The overall PIDs in the USA studied and modelled by Kitov (2005b) demonstrate a very stable structure of the society in respect to distribution of the total income. However, the observed PIDs change very fast from one age group to another. The personal income data sets with high age resolution for years between 1994 and 2002 include information on the number of people in 5 year wide age intervals starting with 15 years of age. The first interval is twice as long, however, and spans the age interval between 15 and 24 years. This widening reflects enumeration problems in this age group, where standard definitions of income (as listed in the technical paper 63RV of the US Census Bureau (2002b)) do not cover all potential income sources as, for example, interfamily money transfer.

Figures 1 and 2 display the PIDs measured in 1998 in various age groups in absolute values and normalized to the total population above 15 years of age respectively. In the youngest age group, the distribution is close to an exponential one with strong variations observed at

incomes above \$50K. This is a result of the Current Population Survey (CPS) and Annual Social and Economic Supplement (ASEC) to the CPS procedures covering only approximately 100,000 households (The US Census Bureau 2002a). When the PID in this relatively narrow age group drops by three orders of magnitude, this coverage can not provide adequate estimates. At higher income, there are no counted people at all. It is expressed by gaps in the distribution curves or zero values in corresponding tables. Thus, one should always keep in mind that the high income tail of the age-dependent PIDs are not reliable and the strong variations have an artificial character representing no actual distribution. (These artificial gaps are also partly responsible for the underestimated average income in the youngest age group as will be discussed in the next paper.)

With increasing age or work experience (equal to the actual age of a person minus the work starting age of 15 years), the PIDs obtain a slowly extending quasi-flat part and a quasi-exponential decreasing part. The PID for the age group between 70 and 74 years is characterized as a whole by an exponential roll-off similar to that for the youngest group. This indicates that old people in the USA lose income very fast with time and not many of them can retain the same income as they had before: compare the oldest group with the group between 60 and 64 years of age. Figure 2 illustrates this process in relative units. The curve corresponding to the oldest reported age group lies below the curve for the age group between 25 and 29 years for incomes higher than \$30K and well above this curve for incomes below \$20K. The middle age groups are characterized by very closely bunched population density-income distributions. Thus, the observed PIDs for a single calendar year (1998) uncover a complex character of their evolution with age. The PIDs for other years of the studied interval are similar and show the same principal features.

Of the same principal importance is the evolution of the PIDs with time observed in the same age groups. Figure 3 shows two PIDs (current dollars) in the youngest age group for calendar years 1994 and 2002. An exponential regression gives a negative index, increasing from -0.125 in 1994 to -0.095 in 2002. The ratio of the indices is 1.32. This value is very close to that observed for per capita nominal GDP growth from 1994 to 2002, equal to 1.34 (Bureau of Economic Analysis 2005). Thus, an adjustment for the per capita nominal GDP growth like the one applied to the overall distribution by Kitov (2005b) would completely merge the curves.

Figures 4 and 5 show results of an exponential regression analysis of the second (or decreasing) PID part (the quasi-flat parts are not shown) for the age groups from 24 to 29 years and from 45 to 49 years. Figure 6 summarizes results of the exponential regression study depicting the indices obtained for calendar years 1994 and 2002 in various age groups. The index curves decrease with age according to a power law. Both calendar years are characterized by the same exponent of -0.5. Thus, one can assume that the ratio of indices for all the ages is the same and equal to 1.33, i.e. the ratio of the power law coefficients,  $0.2864/0.2011=1.33$ , and that all the PIDs return to that for 1994 if corrected for the per capita nominal GDP growth. This assumption may be not so accurate, however, for the age groups around 50 years.

There are four principal features of the original PIDs dependence on work experience we would like to stress. The first is the exponential character of the distribution in the youngest age group over the whole reported income range. The second is the development of the quasi-constant part of the normalized distributions extending from a zero income to approximately \$30K in the age groups above 25 years. The third important phenomenon is a fast decrease of the distribution for the oldest age group. The distribution in this group shows a character similar to the distribution in the youngest age group. One can assume that in some older age group, above 75 years of age, the distribution is the same as in the youngest group. And the fourth observation is the synchronized development of the quasi-exponential roll-off parts of the distributions over the years. Despite the change of the exponential indices with work experience, their ratio is almost the same for the 1994 and 2002 PIDs.

The adjustment for the per capita nominal GDP growth effectively returns the overall PID to the original one, as shown by Kitov (2005b). The same procedure is applied to the PIDs in the studied age groups. Some results of the data processing are presented in Figures 7 through 9. In the youngest age group, the adjustment results in the same pattern as observed for the overall distribution. The only difference is a higher scattering at higher incomes related to the lack of reliable data in this income range, as discussed above. The largest differences between the adjusted PID are observed for the age groups from 45 to 49 years and from 50 to 54 years (Figures 8 and 9, respectively). These groups are near the age where the mean personal income distribution over work experience changes from growth to fall. This critical age also changes with time as the square root of the per capita real GDP growth (Kitov 2005a). So, one can expect

the largest disturbance in the PID near this age, which is currently between 50 and 55 years of age in the USA.

## **2. Modelling personal income distribution in various age groups**

The personal income distributions in the age groups published by the U.S. Census Bureau are calculated in the course of the overall prediction by using the macroeconomic model developed by Kitov (2005a). Since the model calculates each and every individual income, it is a very simple procedure to calculate the number of people in any given age and/or income interval. Actually, the model takes into account only 841 distinct individual income histories for every single year of age, but corrects for the total number of people having the same age. Thus, people of the same age are effectively divided into 841 equal groups, and size of these groups changes with age.

Briefly the model presented by Kitov (2005a) is as follows. A personal income depends on two principal parameters. The first one is the economic growth expressed in per capita (real or nominal) GDP growth rate. The second parameter is a product of the personal capability to earn money and the size of earning means. Having these parameters, one can predict any personal income trajectory by using relationships (2) and (3) from (Kitov 2005a). Each trajectory has two branches: increasing and decreasing. The former is described by a function of type  $(1 - \exp(-\alpha t))$ , where  $\alpha$  is a small parameter dependent on the personal size of earning means, and  $t$  is the personal work experience. The second branch is exactly described by a function  $\exp(-\alpha_1 t)$ , where  $\alpha_1$  is also a small parameter, which in general case is not equal to  $\alpha$ . The increasing branch transits into the decreasing one at some critical work experience  $T_{cr}$ . The value of  $T_{cr}$  evolves in time as the square root of the per capita *real* GDP. In 2002, the  $T_{cr}$  was of about 40 years in the USA.

In absolute terms, both factors of the product representing the capacity to earn money grow proportionally to the square root of the observed per capita (real or nominal depending on whether the data presented in real or nominal dollars) GDP making the product to be proportional to the per capita GDP growth itself. This growth of the absolute value of the personal income with calendar year corresponds to a drift coordinated with the overall evolution of the economy. Thus, each and every personal income is assumed to be growing proportionally to the observed GDP growth rate. At the same time, each personal income evolves relative to

other personal incomes as the person gets older. (One can consider two time scales: a global time scale for the economy as a whole – calendar years; and a local scale for the person – age). In the beginning, this relative growth is fast and then it is slowing down to reach a zero growth rate at  $T_{cr}$ . If there is no economic growth in the country ( $T_{cr}$  is constant), the pattern of the relative growth will repeat itself: the personal income trajectories are all the same and do not depend on the starting year of work. Any economic growth, whether positive or negative, leads to the change of the size of earning means and  $T_{cr}$  as well. The change distorts the personal income trajectories and the overall pattern of the relative personal incomes' growth.

For people of the same age (for the sake of simplicity we assume that all the people of the same age have the same birthday and start day of work), there are only 841 possible personal income trajectories. These trajectories correspond to 841 distinct combinations of the individual capability to earn money and the size of earning means available according to the even distribution of these two factors as integer numbers between 2 and 30 ( $841=29 \times 29$ ).

Everybody can choose one of the trajectories and follow it in time. The person can also jump from one trajectory to another. The latter forces somebody who had before the new trajectory of the first person to drop to the left trajectory, i.e. the trajectories can be only swapped not created. Thus, the relative distribution of the personal incomes is the same for people of the same age. Ratio of personal incomes of two persons with the same capacity to earn money and the same size of earning means but of different age depends on the observed economic growth during the period between the calendar years they started their job careers.

The model's defining parameters  $T_{cr}$  and  $\alpha$  depend on time in a developing economy (Kitov 2005a). A series of calculations has been conducted for various time intervals between 1950 and 2002 in order to obtain the best estimates of the defining parameters (Kitov 2005b). The best fit parameters for 1960 (the start year of the age dependent PID study) are  $T_{cr}=26.5$  years and  $\alpha=0.087$ , respectively. These values can be reduced to any year other than 1960 by using the per capita real GDP growth rate.

Figures 10 through 13 show the observed and predicted distributions in the age groups from 15 to 24 years, from 25 to 29 years, from 30 to 34 years, and from 60 to 64 years in 1998. This year has an advantage of a very simple conversion factor between the model dimensionless units and the observed currency units (dollars) equal to 10,000. This means that 1.0 model unit costs \$10,000 in 1998 and scaling of the model results to those observed is straightforward. As



has been already shown, the other middle age groups have distributions very similar to those for the age groups from 30 to 34 years and from 60 to 64 years.

In the youngest age group, the distributions diverge almost everywhere. The predicted distribution lies above the actual one. One can explain this observation by the factors discussed above: the low resolution in this group, the undercounting at higher incomes and the absence of adequate income sources in the Census Bureau's questionnaire. So, this deviation is expectable and should be resolved somehow in further income surveys, i.e. this deviation is induced by some deficiencies in the current design and methodology of the surveys not by the model. Both distributions are characterized, however, by the same indices of the exponential decay from the zero income bin to the income bin starting at \$40K. The observed distribution has already an emergent part characterized by the Pareto distribution (above \$48K) and also has a large number of people with very low income. The latter observation is consistent with the predicted behaviour. Again, the accuracy of the enumeration at lower incomes is under doubt due to the limited questionnaire.

In the middle age groups, evolution of the predicted and observed distribution is described with reasonable accuracy. Here we meet again the low income counting problem. In the oldest age group among the ones presented here (from 60 to 64 years of age), the theoretical curve fits the measured one with an excellent precision. As in the overall PID (Kitov 2005b), one can distinguish a sub-critical zone, a zone where the theoretical distribution is above the observed and a zone of an opposite behaviour. The same age group has been used for calibration of the earning capabilities (Kitov 2005a).

The actual PIDs lack resolution in the most important age group – from 15 to 24 years of age where dynamics of the income evolution is the fastest. The observed PID for this age interval actually consists of ten very different single year PIDs. Figure 14 displays the predicted evolution of the single year of age PIDs for 1998 starting from 3 years of work experience, i.e. decomposition of the aggregated distribution into the single year of age distributions. During the first two years of work the predicted incomes are concentrated in the lowermost income bins. The predicted evolution reveals some important features of the PIDs. During the first eight years of work nobody was able to reach the Pareto distribution income range starting from \$48K (the Pareto distribution threshold). This can be the actual reason for the U.S. Census Bureau to aggregate all the personal incomes between 15 years and 24 years of age. Otherwise, there is

almost nobody filling the high income bins in the Pareto distribution range. (There are actually quite a few people starting their job career before reaching the age of 15 years. There is no official statistics for these people, however, and their overall impact is negligible in the course of the PID evolution because they affect only the youngest age group with the lowermost incomes. With age, all these differences in the start year of work are smoothed.)

In addition to a higher age resolution, the predicted PID can be extrapolated years back and ahead. Figure 15 and 16 compare the predicted PID in 5 year wide work experience intervals for calendar years 1980 and 2002. One can observe that the PIDs span very different ranges of income despite the same virtual procedure of counting in \$2500 wide intervals. In 1980, the predicted distributions lie inside the income interval from \$0 to \$35K. The Pareto distribution threshold was of about \$20K according to the nominal per capita GDP increase by 2.9 times from 1980 to 2002. (An actual distribution, if measured, would be characterized by the Pareto law above \$20K and would be extended to very high incomes well above the model predicted \$35K.) In 2002, the Pareto threshold was of about \$58K and the predicted distributions occupy the whole income range of the actual survey – from \$0 to \$100K. One of the effects of the observed PID stretching with time is that relatively lower number of people occur in the predefined income intervals. This effect results in the observed oscillations in the PIDs. The PIDs in 1980 look much smoother than in 2002 due to smoothing effects of the large numbers. Oscillations in the predicted PID are only induced by the discrete distribution of the capacity to earn money, which was revealed by the calibration procedure described by Kitov (2005a).

### **3. Discussion and conclusions**

The personal income distribution in the USA was modelled for the period between 1994 and 2002 in the second paper of the series (Kitov 2005b) and found an excellent match between the predicted and measured PIDs. This fact allowed us to assume that the model is adequate and to calibrate the model as well. Defining parameters of the model obtained and described in the first two papers occurred to be accurate in prediction of the PIDs' evolution in time. The prediction of the evolution was almost solely based on the assumption that the critical time,  $T_{cr}$ , and  $\alpha$  evolve as the square root of the per capita real GDP.

This paper addresses a more complicated problem of age structure of the observed PIDs. Complexity of the problem is obvious due to cardinal changes in the PID shape with age. The model, however, meets this new challenge and accurately predicts not only the overall behaviour of the age-dependent PIDs, but also some fine details as presented above. Moreover, the model unveils some shortcomings of the income survey methodology and design, which lead to degradation of the observations' accuracy with time.

As discussed above, one needs only to predict the PID in the lower income zone below the Pareto threshold. The distribution above the threshold is controlled by some unknown processes, which are called self-organized criticality. The name used in natural sciences. The processes may differ from one discipline to another, but have the same result – a power law distribution of sizes. Thus, the only value one has to predict to model the Pareto distribution is the number of people who have reached the threshold depending on age and calendar year. Results of this study are presented in (Kitov 2005c).

The data on the age-dependent PIDs are obtained from the U.S. Census Bureau web-site. Two important features of the data should be mentioned here. Definition of income adopted by the Bureau in its questionnaire is very limited and does not comply with broader definitions based on consumption or expenditures. This kind definitions better present the production side of the income definition adopted in the model. The main assumption of the model is that personal income is exactly equal to the total price of the produced goods and services. This assumption effectively balances net expenditures and incomes in the society. Difference between the two definitions under discussion is the largest in the lower income zone where pure money earnings sometimes are not the major source of income. Interfamily money transfer also can be of importance in this zone. The difference is clear when we compare the observed and predicted distributions in this zone. The majority is concentrated in the first income bin from \$0 to \$2500, where also people with total loss are placed. There are some doubts, however, that a person without any income can survive for a long period of time. If to consider the real expenditures of the person to stay alive as his/her income we effectively have our definition of the personal income. At higher incomes, the difference is obviously lower because principal sources of income have here some monetary form. It is worth noting also that the net income in the poorest group is less than that in the middle and higher income groups and plays only a marginal role in the total economic development.

The second feature of the data is related to the coverage of the age and income intervals. With the PIDs stretching with time over wider and wider income range, smaller and smaller number of people are counted in the designed income bins. In the youngest age group, the number of people changes by almost four orders of magnitude, i.e. if there are 10,000 people in the lowermost income interval, less than ten people are present in the highest measured income bin (See Fig. 1). In some income bins there is nobody counted in. This observation does not present the real personal income distribution, but reveals only increasing problems in the current survey design. There are two ways out: not to publish the data characterized by very high uncertainty or to increase the population coverage (number of households) to make the data accurate and representative.

The presented model is a good basis for the development of a new methodology for the income distribution measurements and definitions of economic equality/inequality. Current inequality measures like Gini coefficient and Lorenz curve usually say next to nothing about the causes (driving forces) of the observed income distribution, are of artificial character, and are age independent as well. The model predicts the evolution of the PID with economic growth and reveals important future changes. For example, the overall economic growth results in longer time necessary to reach higher income. This makes young people relatively poorer - a trend which is observed in the USA. Further increase of  $T_{cr}$  and decrease of  $\alpha$  will accelerate this process in near future. Also, people reach the peak personal income later in age, but before the retirement age: the former approaching the latter with the overall economic growth.

### **Acknowledgements**

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## Figures

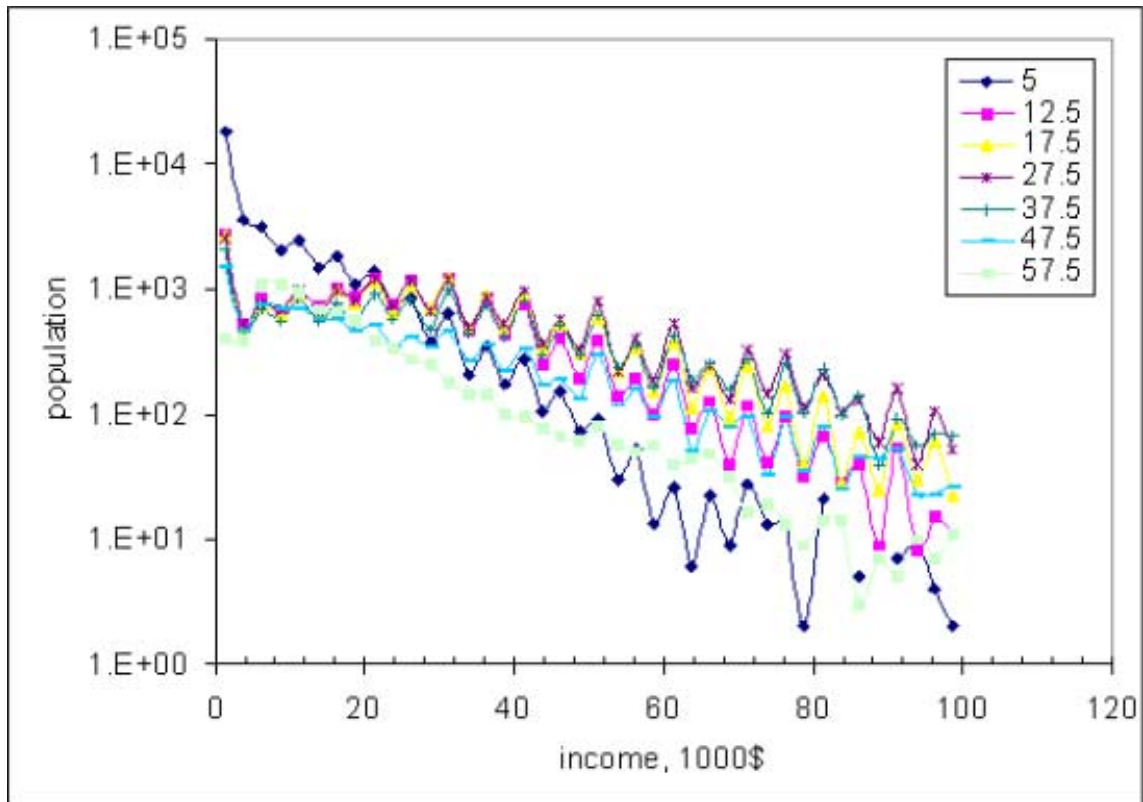


Fig. 1. Personal income distribution in age groups. In the first age group (5) - from 0 to 9 years of work experience, an almost exponential decrease in observed.

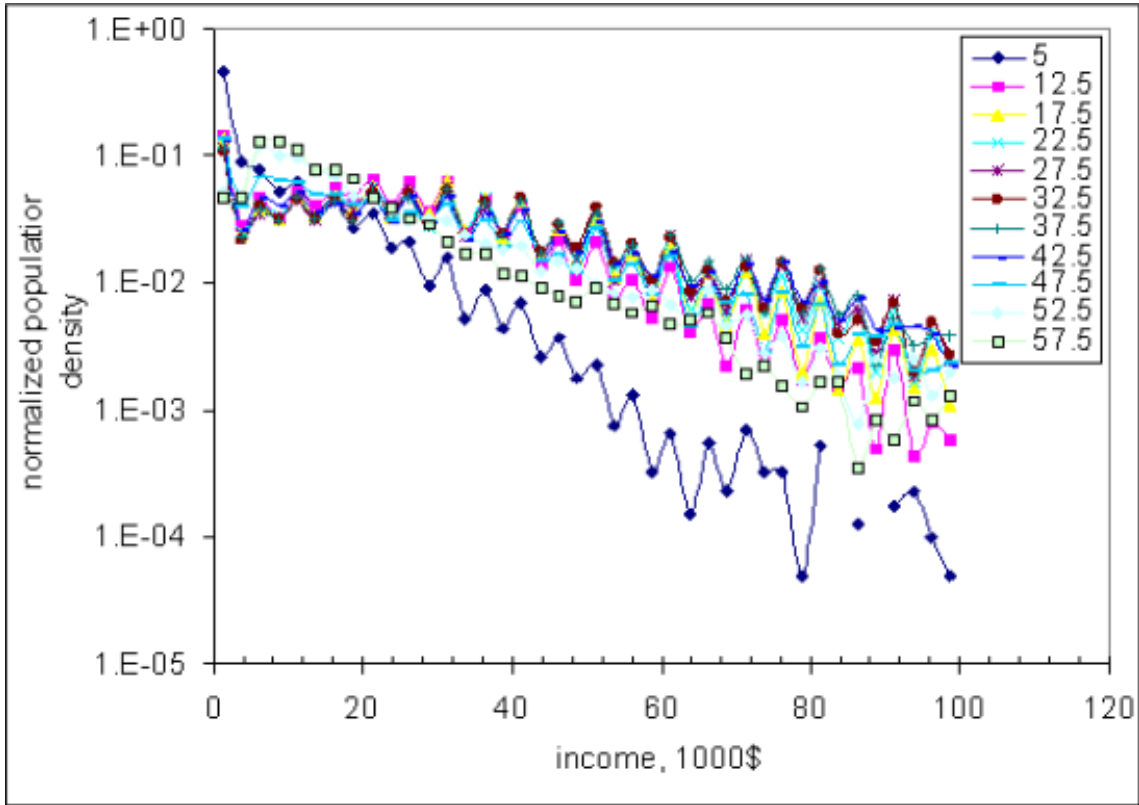


Fig. 2. Population density vs. personal income in age groups from 15 to 24 years (5 – central point of corresponding work experience interval from 0 to 9 years), from 25 to 29 years (12.5), ..., from 70 to 74 years (57.5). In the first age group an almost exponential decrease in observed.

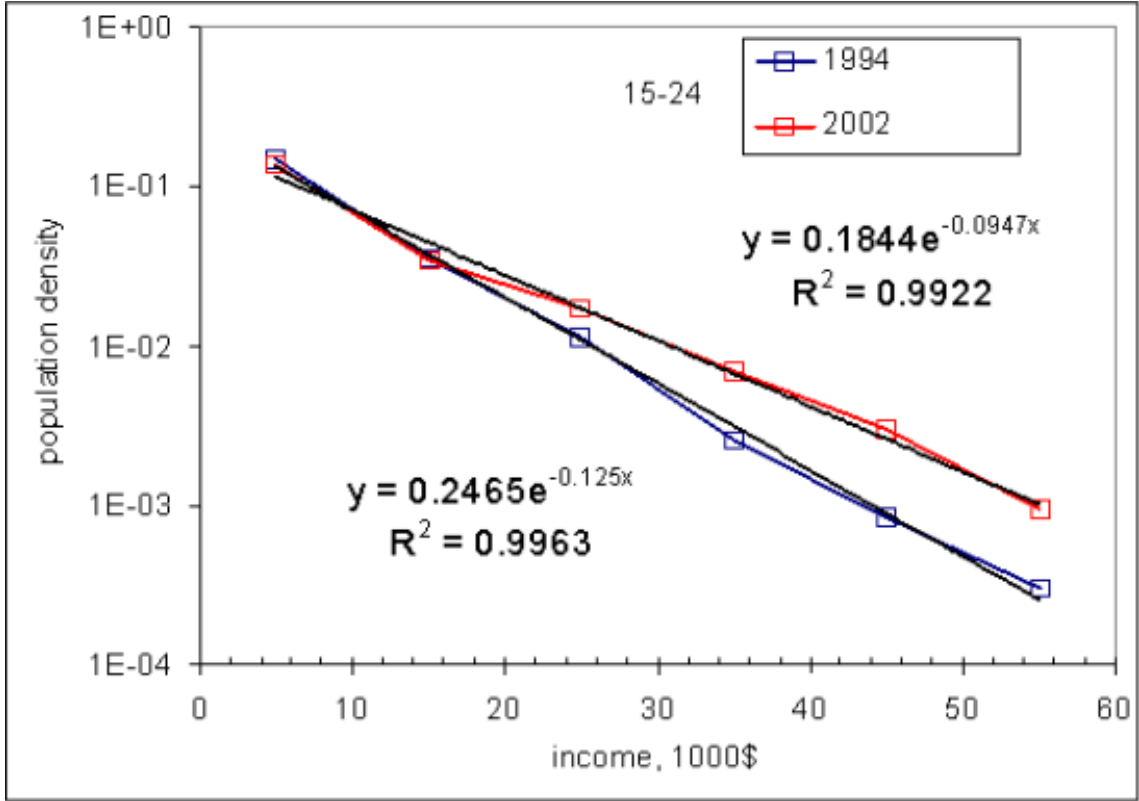


Fig. 3. Population density vs. personal income in the age group from 15 to 24 years. The index of exponential regression function increases with time from -0.125 to -0.095.



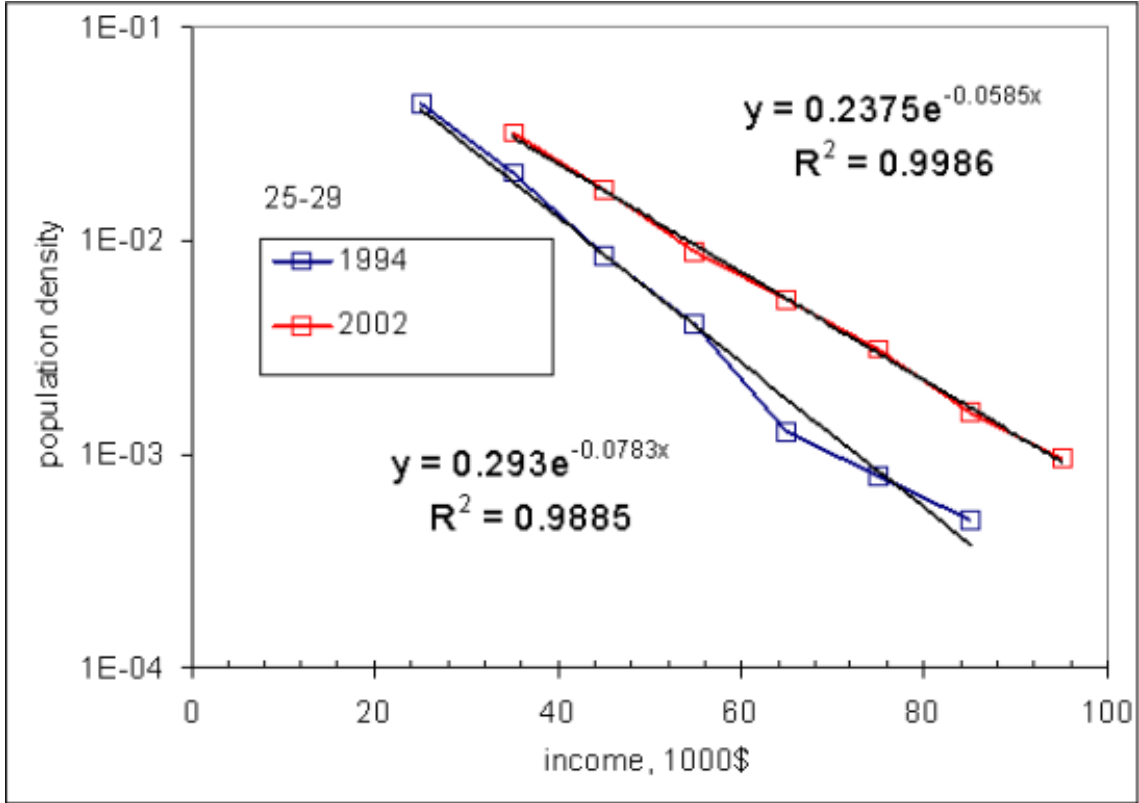


Fig. 4. Exponential decrease of population density function with increasing personal income in the age group from 25 to 29 years. Flat portion of the distribution is skipped. The index of the exponential regression function increases with time from -0.0783 to -0.0585.

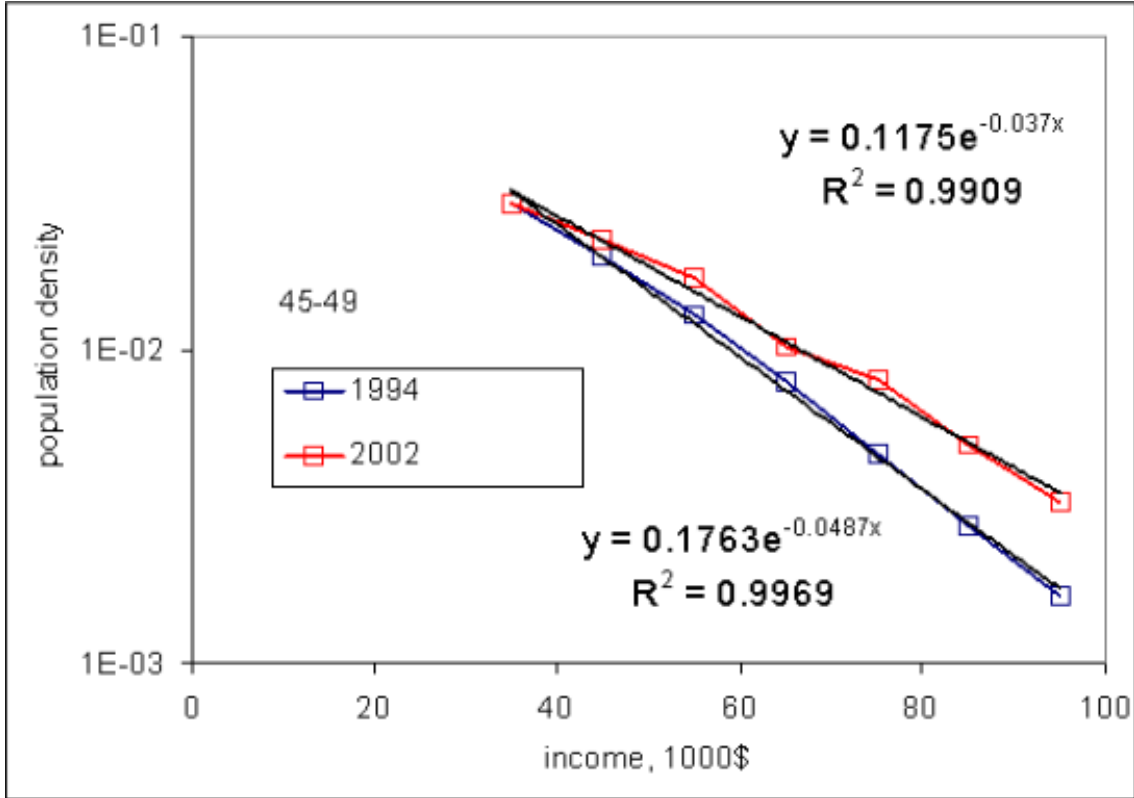


Fig. 5. Exponential decrease of population density function with increasing personal income in the age group from 45 to 49 years. The flat portion of the distribution is skipped. The index of the exponential regression function increases with time from -0.049 to -0.037.

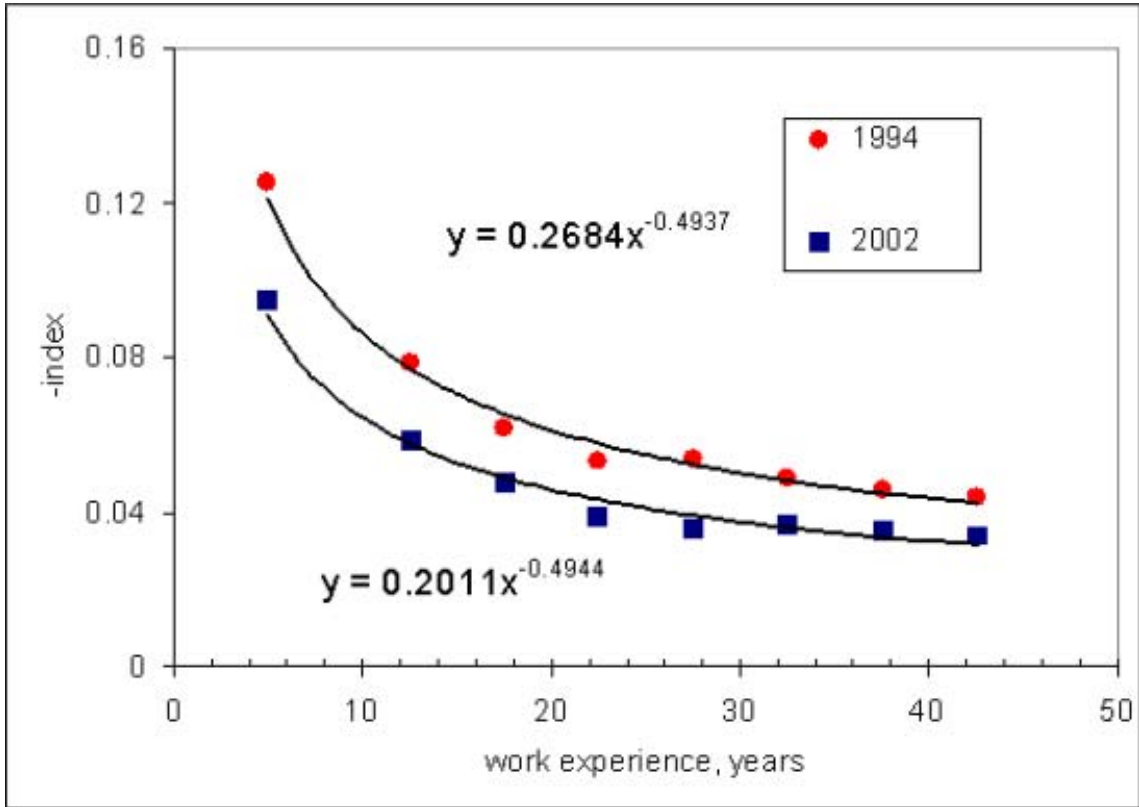


Fig. 6. Index of the exponential regression function vs. work experience for years of 1994 and 2002. The power law regression of the indexes gives an exponent near -0.5 in both cases. The ratio of 1994 and 2002 indexes is 1.33.

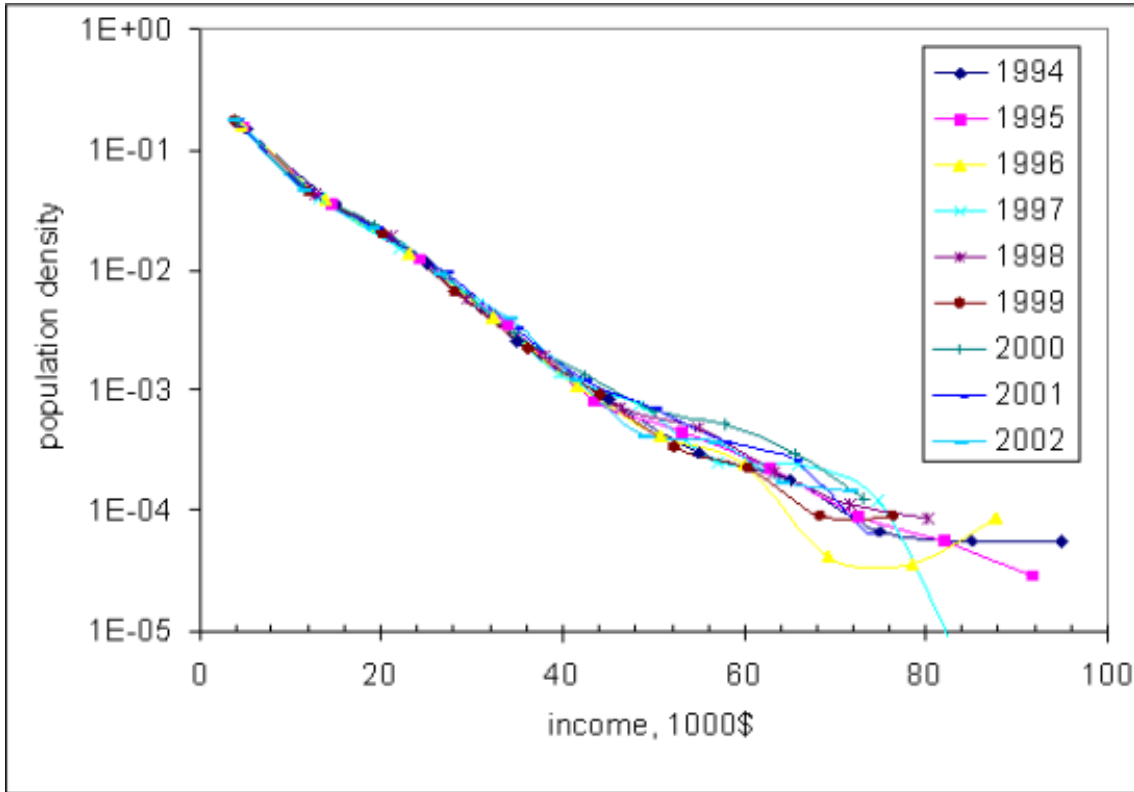


Fig.7. Population density distribution in the age group from 15 to 24 years adjusted for the per capita nominal GDP growth. A strong scattering at higher incomes is induced by lack or resolution power of the current ASEC due to undercoverage of the population. Population density drops by three orders of magnitude with income increase from \$5,000 to \$45,000.

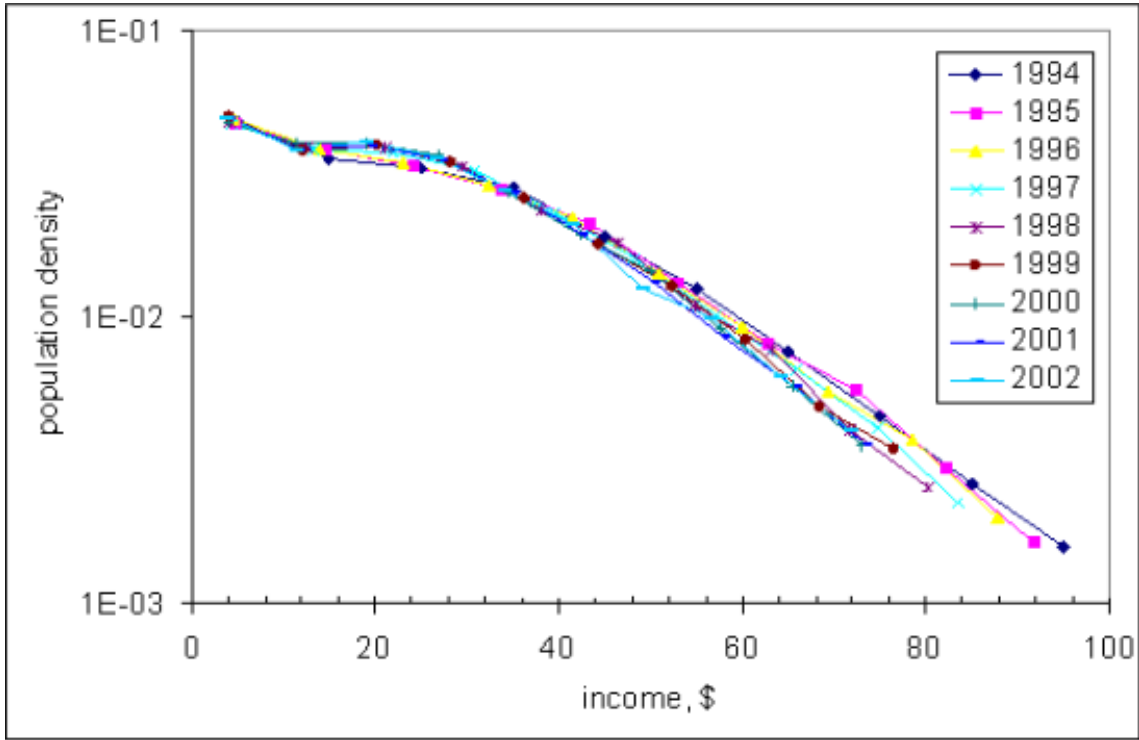


Fig.8. Population density distribution in the age group from 45 to 49 years adjusted for the per capita nominal GDP growth.

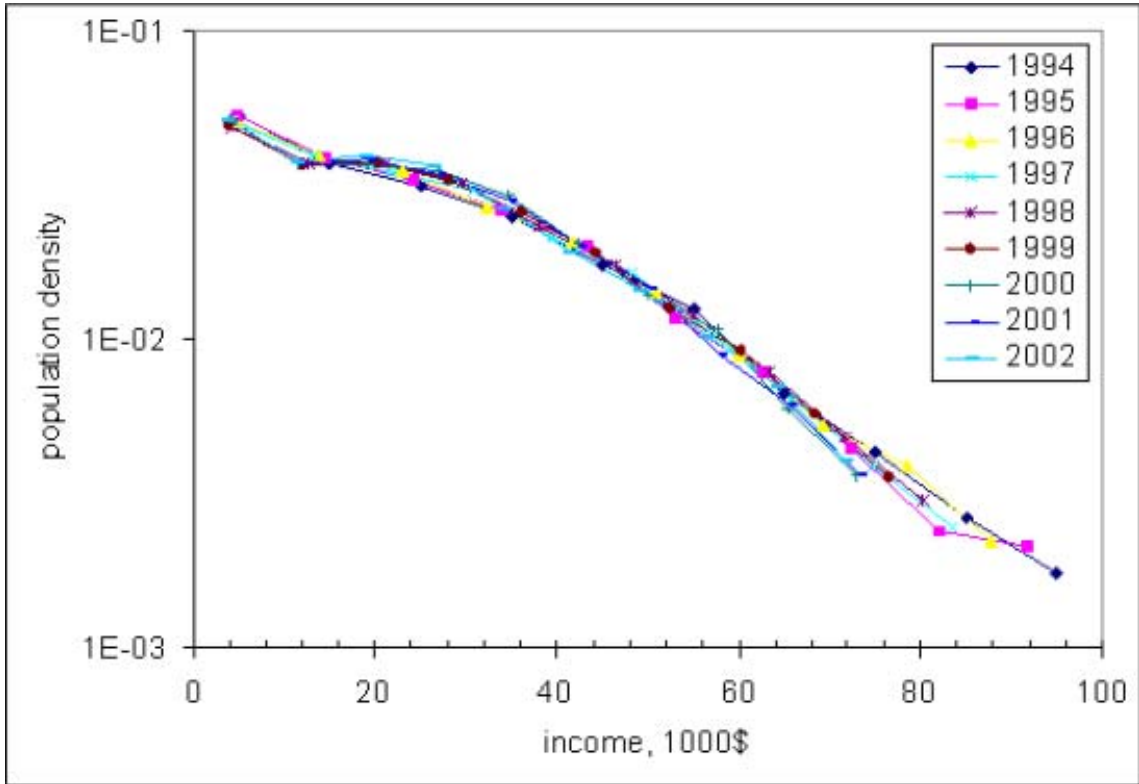


Fig.9. Population density distribution in the age group from 50 to 54 years adjusted for the per capita nominal GDP growth.

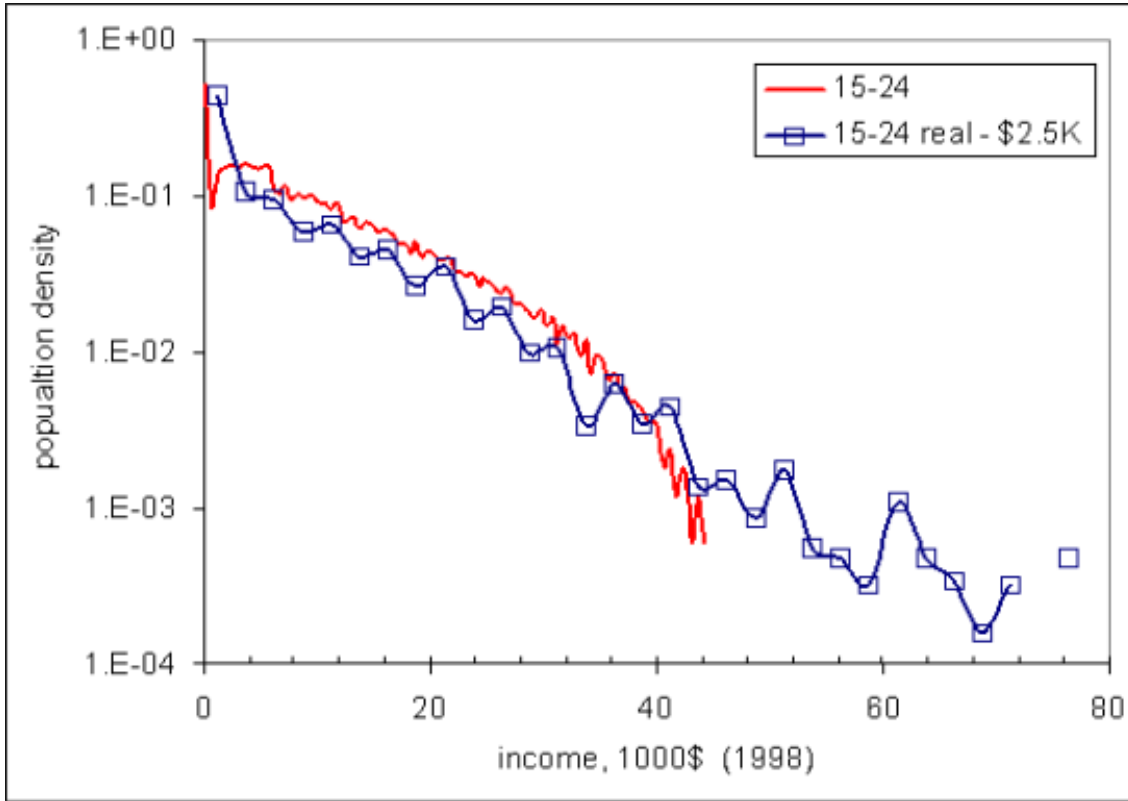


Fig. 10. Comparison of predicted and actual personal income distributions in the age group from 15 to 24 years. The actual distribution is characterized by an emergent Pareto part and lies almost everywhere below the predicted curve. The difference is associated with a poor procedure of the income survey which does not take into account money redistribution.

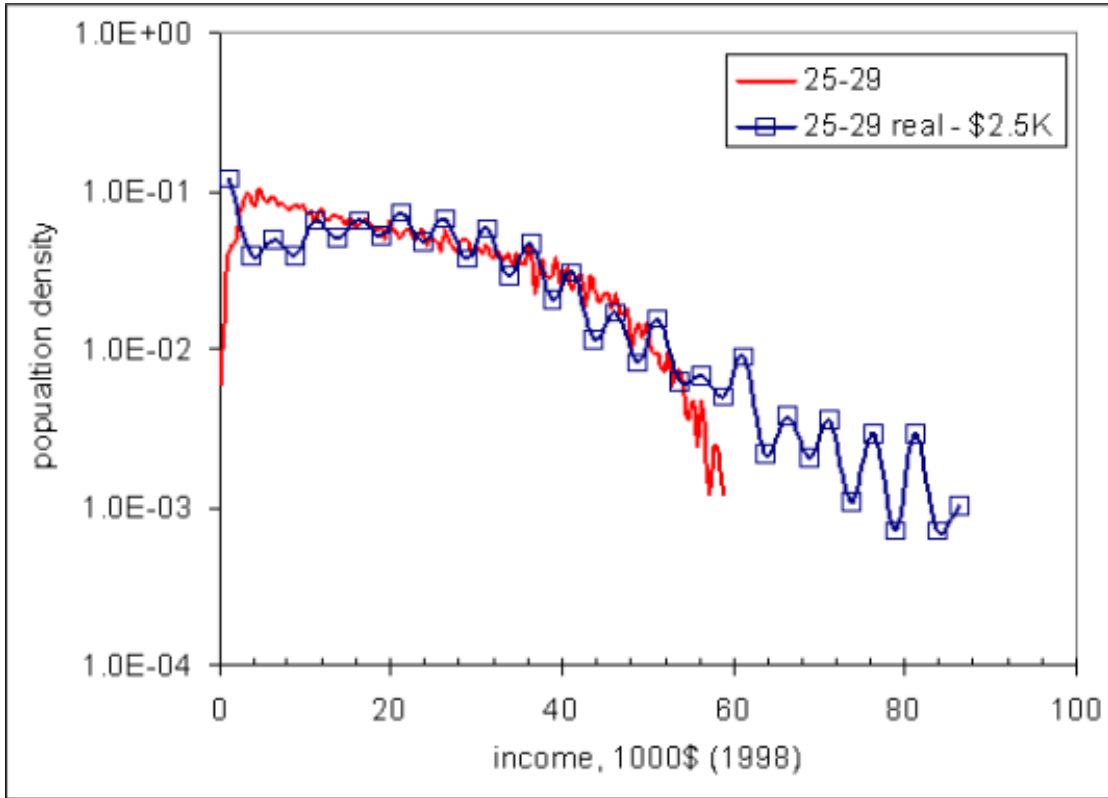


Fig. 11. Comparison of predicted and actual personal income distributions in the age group from 25 to 29 years. The actual distribution is characterized by an elaborated Pareto part. The difference between the curves at low incomes is associated with a poor procedure of the income survey.



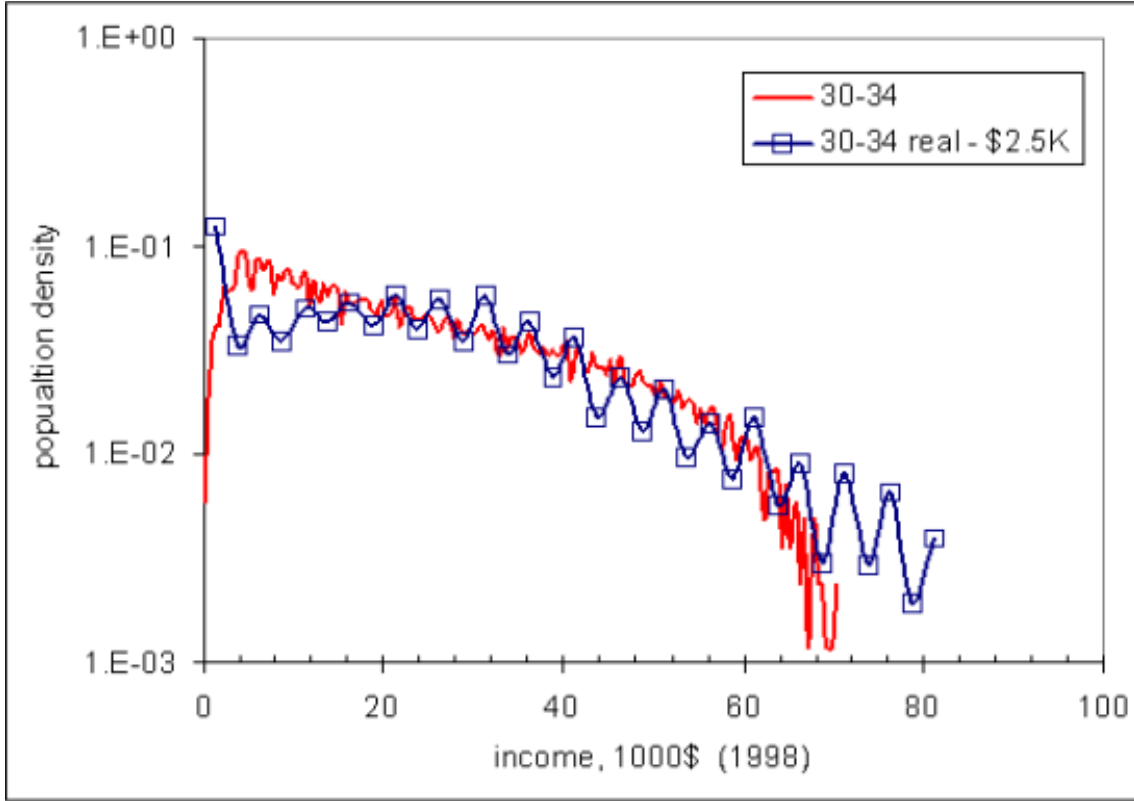


Fig. 12. Comparison of predicted and actual personal income distributions in the age group from 30 to 34 years. The difference between the curves at low incomes is associated with poor procedure of the income survey.

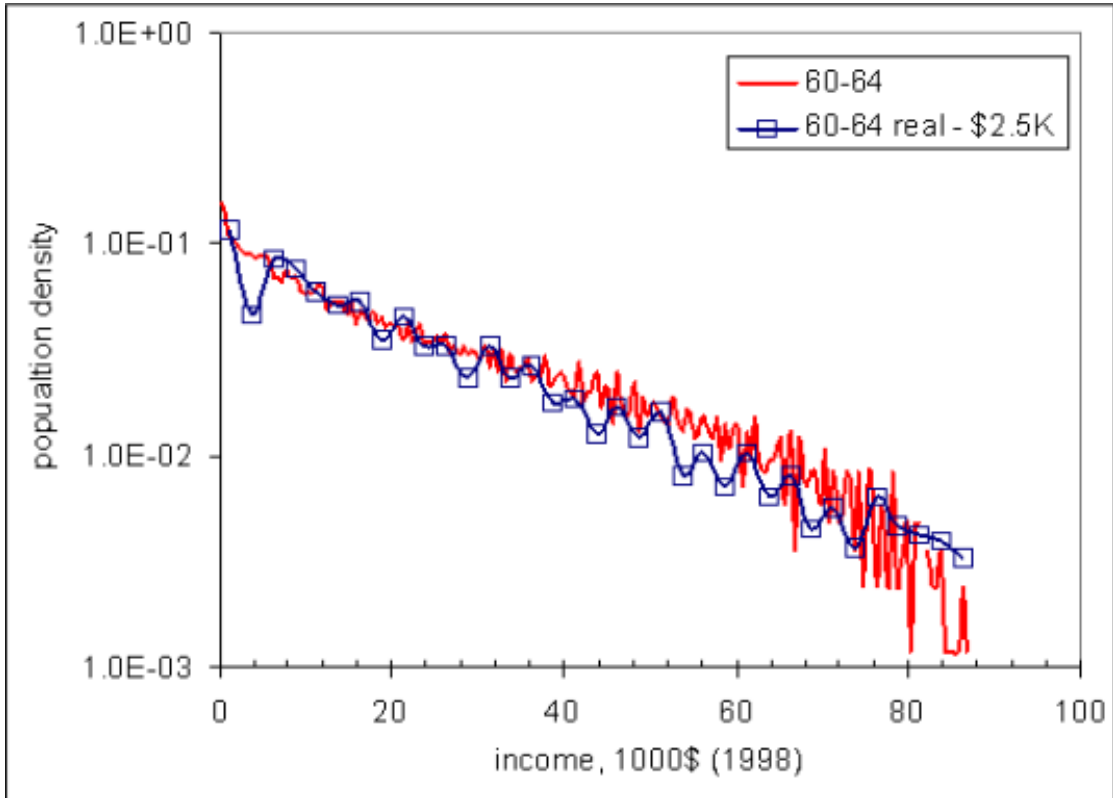


Fig. 13. Comparison of predicted and actual personal income distributions in the age group from 60 to 64 years.

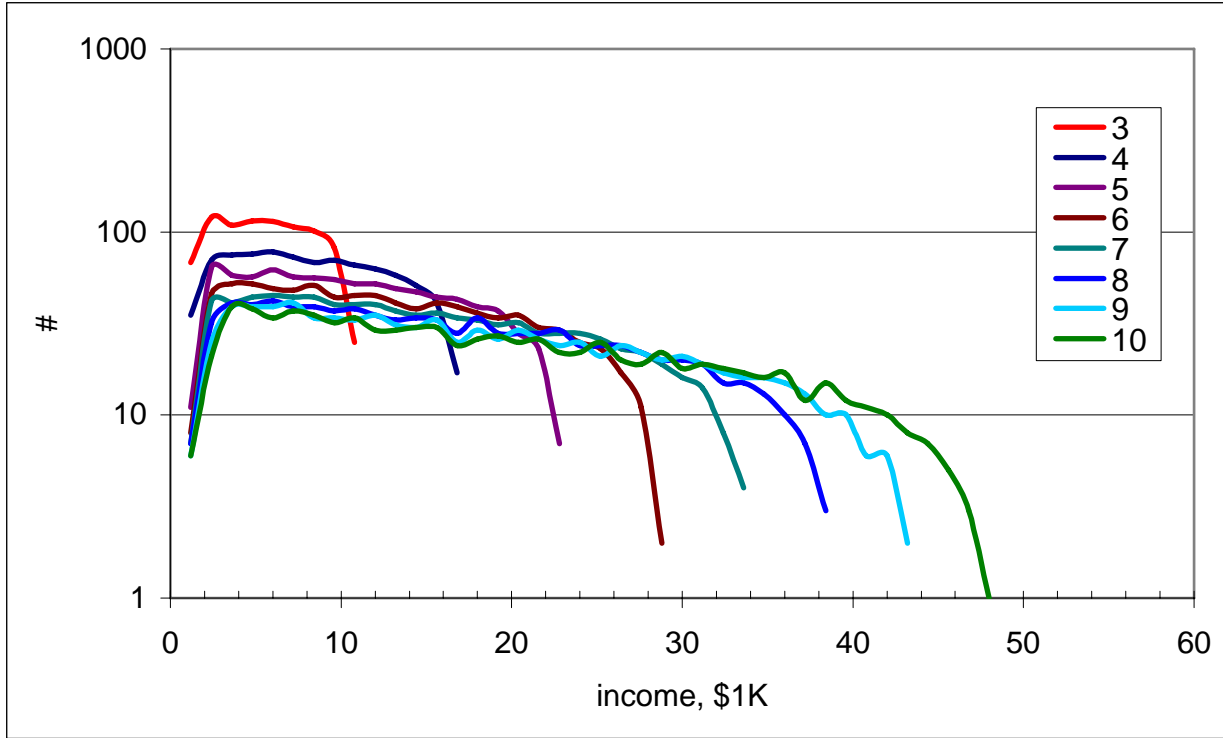


Fig. 14. Evolution of the predicted personal income distribution (absolute number of people) in single year of age intervals between 3 years and 10 years of work experience.

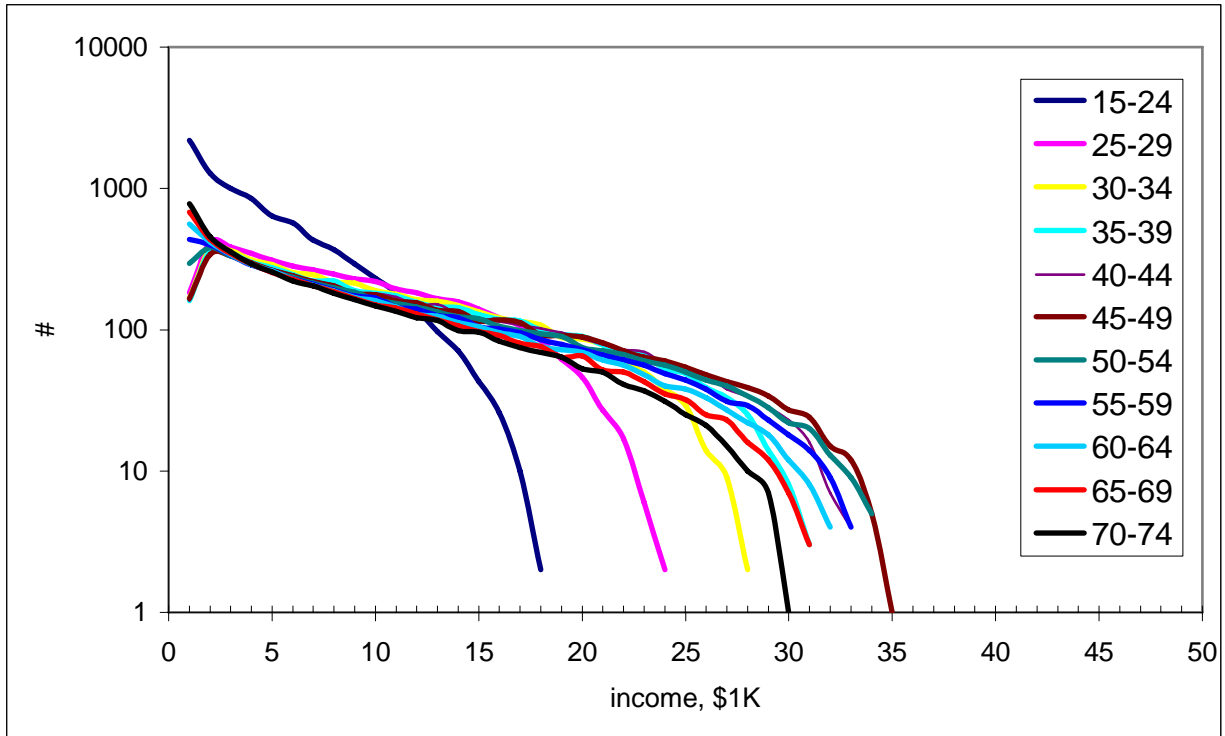


Fig. 15 Evolution of the predicted personal income distribution (absolute number of people) in 1980.

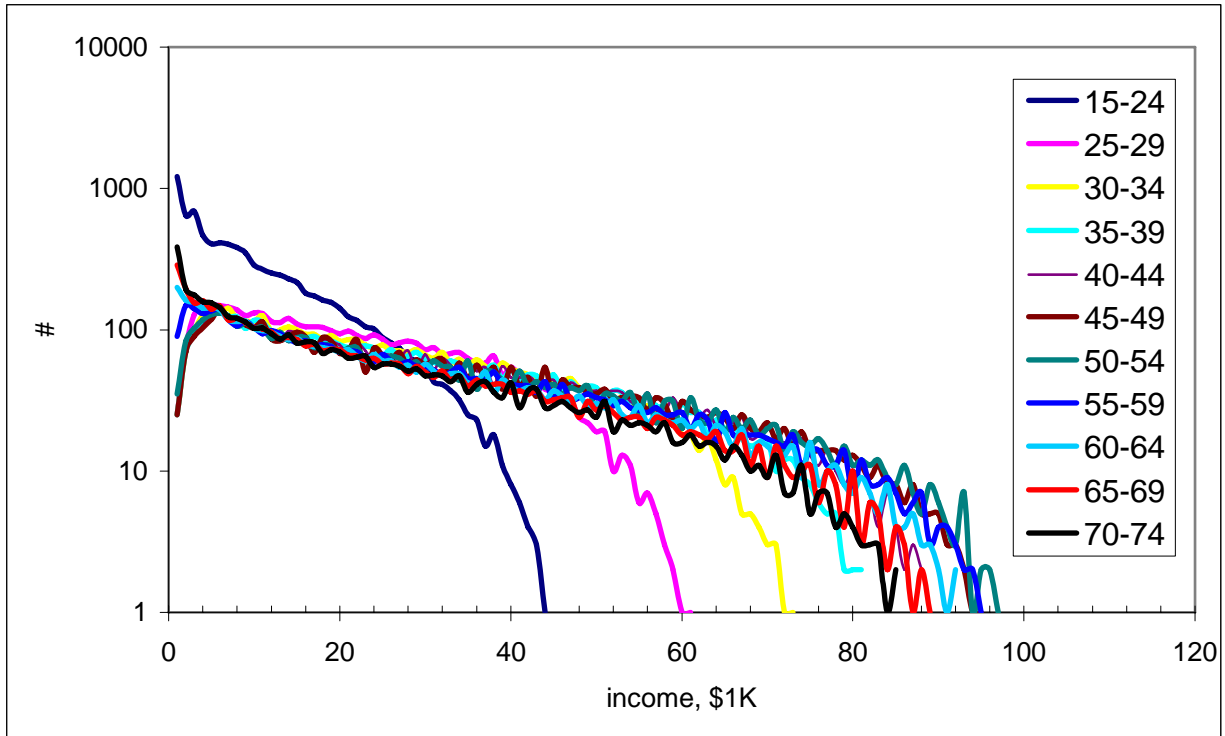


Fig. 16 Evolution of the predicted personal income distribution (absolute number of people) in 2002.